



ELSEVIER

Available online at www.sciencedirect.com

ScienceDirect

journal homepage: www.elsevier.com/locate/jval

The Influence of Genotype Information on Psychiatrists' Treatment Recommendations: More Experienced Clinicians Know Better What to Ignore

Alan J. McMichael, BSc^{1,*}, Marco Boeri, PhD^{2,3,4}, Jonathan J. Rolison, PhD⁵, Joe Kane, PhD⁶, Francis A. O'Neill, MD¹, Ric Scarpa, PhD^{7,8,9}, Frank Kee, MD^{1,4}

¹Centre for Public Health, Queen's University Belfast, Royal Victoria Hospital, Belfast, UK; ²Health Preference Assessment, RTI Health Solutions, Research Triangle Park, NC, USA; ³Gibson Institute, School of Biological Sciences, Queen's University Belfast, Belfast, UK; ⁴UKCRC Centre of Excellence for Public Health Research (NI), Queen's University Belfast, Belfast, UK; ⁵Department of Psychology, University of Essex, Essex, UK; ⁶Institute of Neuroscience, Newcastle University, Newcastle upon Tyne, UK; ⁷Durham University Business School, Durham, UK; ⁸University of Waikato, Hamilton, New Zealand; ⁹University of Verona, Verona, Italy

ABSTRACT

Background: This study applies attribute nonattendance to medical decision making. We aimed to demonstrate how this type of analysis can be used in medical decision making to assess whether psychiatrists were influenced in their treatment recommendations by information on the genotype of a patient, despite knowing the patient's response to treatment as measured by the Positive and Negative Syndrome Scale. A patient's genetic information may be used to predict their response to therapy; such information, however, becomes redundant, and should not influence decisions, once a clinician knows the patient's actual response to treatment. **Methods:** Sixty-seven psychiatrists were presented with patients' pre- or post-treatment scores on the Positive and Negative Syndrome Scale for two hypothetical treatments for schizophrenia. Psychiatrists were also informed whether the patient possessed a genotype linked to hyper-responsiveness to one of the treatments, and were asked to recommend one of these two

treatments. Attribute nonattendance assessed whether the information on genotype influenced psychiatrists' treatment recommendations. **Results:** Years of experience predicted whether psychiatrists were influenced by the genetic information. Psychiatrists with 1 year or less of experience had a 46% probability of considering genetic information, whereas psychiatrists with at least 15 years of experience had a lower probability (7%). **Conclusions:** Psychiatrists and other clinicians should be cautious about allowing a patient's genetic information to carry unnecessary weight in their clinical decision making.

Keywords: attribute nonattendance, discrete choice, medical decision making.

Copyright © 2016, International Society for Pharmacoeconomics and Outcomes Research (ISPOR). Published by Elsevier Inc.

Introduction

Clinicians are becoming increasingly aware of how a patient's genotype can influence their response to treatment [1]. Tailoring treatments according to this anticipated response is known as stratified, or personalized, medicine [2]. In psychiatry, some genetic profiles in the population are associated with an increased risk of schizophrenia. Furthermore, some genetic profiles signal higher potential benefits of particular antipsychotic treatments [3,4], suggesting that for some patients psychiatric treatments could, in the future, be tailored to their genetic profile. Nevertheless, whether or how information about a patient's genetic profile influences psychiatrists' treatment recommendations is still unclear.

Genetic information may indicate the potential benefits that a patient could receive from a treatment but is redundant when

the patient's actual response to a treatment is known. Thus, in certain circumstances, genetic information about a patient could bias the psychiatrist's clinical decision making. In particular, clinicians may view treatment outcomes differently when they are aware that the patient possesses a genotype that is indicative of hyper-responsiveness to a treatment. Consequently, when aware of a patient's genetic profile, a clinician may be less or more likely to recommend or continue a treatment even though the treatment may have been shown to be effective in the patient's pre- or post-treatment scores on a given symptom report scale. The potential for genetic information to bias clinical decision making in respect of a patient's treatment is known as *pharmacogenetic exceptionalism* [5]; this may result in an inefficient allocation of resources for public health. This article explores the topic by using a choice-format conjoint analysis (referred to as a discrete-choice

Conflicts of interest: The authors declare that they have no conflicts of interest.

* Address correspondence to: Alan J. McMichael, Centre for Public Health, Queens University Belfast, Royal Victoria Hospital, Belfast BT12 6BA, UK.

E-mail: amcmichael01@qub.ac.uk.

1098-3015/\$36.00 – see front matter Copyright © 2016, International Society for Pharmacoeconomics and Outcomes Research (ISPOR).

Published by Elsevier Inc.

<http://dx.doi.org/10.1016/j.jval.2016.09.2395>

experiment [DCE]) administered to psychiatrists in Northern Ireland.

In the practice of DCEs, respondents are presented with a sequence of choices for alternative options and are asked to select the one they prefer, with each alternative being described by different attributes and attribute levels [6–9]. A recent review showed a substantial increase in the application of DCEs in health economics and medical decision making and a desire to incorporate patients' and doctors' preferences in the study of effectiveness of treatments [10]. Indeed, the Food and Drug Administration recently stated that new cancer treatments must first assess patient preferences before becoming widely available to all patients [11]. The conventional underlying assumption of DCEs is that when choosing between alternatives, respondents rationally consider all the attributes presented and select the alternative that maximizes their utility. Nevertheless, research has seen an increasing focus on decision-making heuristics [12–14]. One particular type of heuristic widely explored by choice modelers in transportation [15–17] and environmental economics [18–20] is attribute nonattendance (ANA). In ANA, respondents may ignore one or more attributes that they believe are not relevant so as to simplify the process of choosing the best alternative [21]. The importance of ANA in modeling respondents' choices and preferences has been highlighted by its influence on both coefficient estimations and welfare analysis [17,22]. Recently, ANA has also been extended to health economics [14,23] in which researchers warn that not accounting for ANA may lead to biased health policies [24]. Within the context of medical decision-making research, however, ANA has not been widely used to assess which attributes (if any) are nonattended [23]. Researchers consider ANA a nonrational heuristic that should be included in the analysis to avoid bias, but should not be included if respondents acted rationally, as assumed by the framework in which DCE operates. This study departs somewhat from this perspective, because ANA is considered the correct heuristic that a clinician should apply because the patient's response to treatment is already known, making the patient's genotype information redundant.

This article's contribution to the literature is twofold. From a methodological viewpoint, ANA is applied in a new, present, and highly relevant context—stratified medicine—tackling the issues of coherence of information assessment in the psychiatrist's treatment selection. The novel methodological aspect here is the use of ANA to improve the understanding of the extent to which medical decision making incorporates irrelevant information. From a clinical perspective, the article aims to contribute to the topical issue of whether genotype information influences the treatment recommendations of psychiatrists when a patient's treatment response (in terms of symptom improvement) is already known to the psychiatrist.

Analytic Framework

The analysis of a DCE is based on the random utility maximization theory [25,26] in which the underlying assumption is that individuals select the alternative that offers them the highest utility. In this context, it is possible to denote with i the treatment that psychiatrist n recommended when considering the vignette t . The utility function that psychiatrists maximize when recommending a treatment can be described by characterizing each vignette using a vector of attributes (X) and a vector of parameters (β) to be estimated as follows:

$$U_{nit} = \beta' X_{nit} + \varepsilon_{nit}, \quad (1)$$

where ε represents the part of the utility function that the researcher cannot observe and is assumed to be an independent

and identically Gumbel-distributed error term. With these definitions and assumptions, it is possible to mathematically specify the choice probability for each psychiatrist n selecting treatments i over j alternatives in the vignette t as a multinomial logit (MNL) selection probability [26,27]:

$$\Pr(\text{nit}) = \frac{\exp(\beta' X_{nit})}{\sum_{j=1}^J \exp(\beta' X_{njt})}. \quad (2)$$

This model is estimated as a benchmark and is the simplest starting point for behavioral analysis. Notwithstanding the importance and practicality of the MNL model results, the MNL has several restrictive assumptions. For example, preferences are homogeneous across respondents and choices are independent from irrelevant alternatives. These assumptions are often considered unrealistic and are likely to bias the results [28]. The mixed logit (MXL) model relaxes the restrictive assumptions underlying the MNL model and accommodates for the possibility that respondents may have different preferences [29]. Furthermore, the model fit to observed data is typically improved when estimating MXL models [30]. The models derived within the general framework of the MXL allow for taste parameters β to vary across respondents and to account for the fact that in the DCE, each respondent is observed across a series of T vignettes and can therefore be represented as a balanced longitudinal panel of responses on experimentally designed choice tasks (vignettes). If the value of β was known for each of the n th respondents, the probability of a sequence of choices would be given by Equation 3:

$$\Pr(y_{Tn} | \beta, X_{nit}) = \prod_{t=1}^T \frac{\exp(\beta' X_{nit})}{\sum_{j=1}^J \exp(\beta' X_{njt})} \quad (3)$$

Because it is impossible to know the value of β with certainty for each respondent, heterogeneity of preferences is estimated by allowing for random variation in β across respondents [7,31]. To address the research question, it is essential to understand whether psychiatrists are influenced by information about a patient's genotype in making their treatment recommendations. Therefore, we were interested in modeling ANA in this context while addressing preference heterogeneity. In this article, ANA was analyzed by means of behavioral latent class (LC) models, which are semiparametric variants of the MNL model. In LC models, it is assumed that each individual respondent can be implicitly sorted into a set of C behaviorally defined classes associated with certain estimated probabilities, with each class characterized by a unique class-specific pattern of ANA embedded in the utility parameters, β_c . With membership to class c , the probability of respondent n 's sequences of choices y_{Tn} over T choice occasions is as follows:

$$\Pr(y_{Tn} | \beta_c, X_{nit}) = \prod_{t=1}^T \frac{[\exp(\beta_c' X_{nit})]}{\sum_{j=1}^J [\exp(\beta_c' X_{njt})]}. \quad (4)$$

Considering that the membership probabilities π for each behavioral LC c are also defined according to an MNL process, we have

$$\pi_c = \frac{\exp(\alpha_c + \gamma_c' z_n)}{\sum_{c=1}^C \exp(\alpha_c + \gamma_c' z_n)}, \quad (5)$$

where z_n is a vector of covariates characterizing respondent n , γ_c is a vector of associated parameters subject to estimation, and α_c is a class-specific constant. In the estimation of LC models, for identification purposes, only $C - 1$ set of coefficients can be independently identified (e.g., for one arbitrary class c , the vector $\langle \alpha_c : \gamma_c = 0 \rangle$).

The probability of a sequence of choices is:

$$\Pr(y_{Tn}|X_{nit}) = \left(\sum_{c=1}^C \pi_c \prod_{t=1}^T \frac{\exp(\beta' x_{nit})}{\sum_{j=1}^J \exp(\beta' x_{njt})} \right). \quad (6)$$

The primary hypothesis of this article was that genotype information might influence some doctors even though this information is redundant. Therefore, this study first focused on a relatively reduced model specification in which ANA affects only one attribute (genotype information). This resulted in a model with only two classes (we ignored ANA on attributes other than genotype information). Given the importance of heterogeneity, the final model accommodated for random variation of preferences across respondents by incorporating a random-parameters logit (RPL) model within each class. The final model estimated was represented as follows:

$$\Pr(y_{Tn}|X_{nit}) = \int \left(\pi \prod_{t=1}^T \frac{\exp(\text{ANA}\beta' x_{nit})}{\sum_{j=1}^J \exp(\text{ANA}\beta' x_{njt})} + (1-\pi) \prod_{t=1}^T \frac{\exp(\beta' x_{nit})}{\sum_{j=1}^J \exp(\beta' x_{njt})} \right) f(\beta) d\beta, \quad (7)$$

where $(\text{ANA}\beta'x)$ denotes the indirect utility of the vignette for those doctors who ignored the information on genotype, whereas those who attended to this information have an indirect utility of $\beta'x$. The probability of nonattending to the information on genotype is represented by π (see Equation 5).

Our second hypothesis was that doctors use other strategies to simplify the decision-making process (because doctors often have to make many decisions very quickly, they might use ANA to simplify their task). Therefore, we extended our behavioral investigation to explore the entire combination of ANA specifications. The combination of ANA behavior across the four attributes, each of which can be attended to or ignored, generated $2^4 = 16$ behavioral classes (Equation 6). The models were estimated using BIOGEME 2.2 (Michel Bierlaire, Switzerland) [32].

Methods

Participants

The sample comprised 67 practicing psychiatrists recruited in Northern Ireland. Respondents were tested during single-session continuous professional development meetings in three hospital trusts. Participants provided their demographic information, whether they had completed their specialist training, and, if so, the number of years of experience in clinical practice and their subspecialty. More than half (59%) were male. Most (64%) had completed their specialist training. The average years of clinical experience in their specialty was 10 years (SD 7.19 years). Ethical permission was granted from the Queens University Belfast Ethics Committee. Each participant also provided informed consent before completing the study.

Vignette Design

Twenty-six vignettes were developed to assess the effect of each attribute on psychiatrists' treatment recommendations for patients with schizophrenia (Fig. 1). Each vignette provided a hypothetical patient's pre- and post-treatment symptom scores on the positive subscale of the Positive and Negative Syndrome Scale (PANSS) for two treatments. The positive subscale of the PANSS consists of seven symptom report items, each rated on a 7-point scale, ranging from "absent" (numerical value = 1) to "extreme" (numerical value = 7). The scores are summed across the seven items to generate a total positive subscale score, ranging from 7 to 49, with higher scores indicative of more extreme symptoms [33]. All vignettes presented a pretreatment score of 42, indicating severe positive symptoms of schizophrenia before treatment [33]. Across the vignettes, the pre- and post-treatment change scores ranged from 3 to 26 points.

Each vignette also identified whether the patient had a genetic biomarker for one of the treatments: participants were told that the genetic biomarker was associated with a 30% increase in the effectiveness of the corresponding treatment. The biomarker was

Patient Information	Treatment A	Treatment B
<p>Displayed to the right are the patients' pre- and post-treatment scores on the positive subscale of the PANSS for Treatment A and Treatment B. The numerical values are provided in parenthesis</p> <p>The shaded horizontal arrows represent the 95% confidence intervals for the patients post-treatment score</p>		
Patient status in respect of <i>hyper-responsiveness</i> genotype to this treatment	Yes	No
Has this patient responded to treatment? (circle as appropriate)	Yes No	Yes No
How confident are you in your judgment (1=not at all confident, 7=very confident)? (circle as appropriate)	1 2 3 4 5 6 7	1 2 3 4 5 6 7
Treatment Information	Treatment A	Treatment B
The cost: acute treatment days in hospital.	30	38
The percentage probability of a 10kg weight gain in the next 6 months following the start of treatment	41%	33%
Based on the information above, which treatment would you recommend? (circle as appropriate)	Treatment A	Treatment B

Fig. 1 – Example vignette. Each treatment showed the full range of scores on the PANSS with arrows showing the patient's pre- and post-treatment scores. Vignettes indicated for which treatment the patient had a hyper-responsiveness genotype. Respondents were asked to state which treatment they would be willing to recommend on the basis of the information available. PANSS, Positive and Negative Syndrome Scale.

present for only one of the two treatments in each vignette. The vignettes also identified two side effects associated with each treatment. One side effect referred to the number of acute treatment days spent in hospital, ranging from 17 to 45 days. The second side effect referred to the likelihood of a 10-kg weight gain over the next 6 months, ranging from 30% to 70%, a common side effect associated with antipsychotic treatment [34–36]. The attributes and levels were based on discussions with two practicing psychiatrists to ensure that the attributes and levels fell within a realistic range that might be experienced in clinical practice. On the basis of the information provided in the vignettes, psychiatrists were asked which treatment they would recommend.

Results

Because we were interested in understanding psychiatrists' preferences for different characteristics of treatments when making a recommendation, we started by modeling their choices by adopting an MNL model and an RPL model to account for heterogeneity in preferences. In both models (Table 1), psychiatrists were significantly more likely to recommend treatments associated with higher post-treatment benefits. As expected, psychiatrists were also significantly less likely to recommend treatments that were associated with more days spent in hospital or a higher likelihood of a 10-kg weight gain. Interestingly, psychiatrists were less likely to recommend treatments for which the patient had a hyper-responsiveness genotype.

Genotype and Its Influence on Psychiatrists' Treatment Recommendations

To test the primary hypothesis related to psychiatrists' attending to the irrelevant information about the patient's genotype, a constrained LC model to control for ANA on only the genotype attribute (as described in Equations 5 and 7) was estimated. This

provided an estimated probability that psychiatrists systematically ignore the information about the patient's genotype. The results of this analysis are reported in the last two columns of Table 1 and suggest that the genotype information did not significantly influence most of the psychiatrists' treatment recommendations. Indeed, across the entire sample of psychiatrists, there was an 84% probability that psychiatrists did not consider the information on patient genotype. Nonetheless, there was a small probability (~16%) that psychiatrists attended to the information. Although this probability is small, it implies that, in some instances, psychiatrists considered the genotype information to be important even though the patient's treatment response on the PANSS was already known to them.

To better characterize psychiatrists who were associated with a positive probability of considering a patient's genotype information when selecting their preferred treatment, we tested the significance of various covariates likely to act as determinants of class membership probability (Equation 6) and found that the number of years of clinical experience was the only significant covariate. Specifically, we found that more experienced psychiatrists were less likely to consider the information on genotype when selecting the treatment to recommend to the patients in the vignette. To be able to expand our discussion on the practical implication of this finding, we simulated posterior probabilities (on the basis of the sequence of choices made by each physician) of being associated with one class or another conditionally to the number of years of experience. The result, as presented in Figure 2, suggests that psychiatrists with less than 1 year of experience had a probability close to 50% of attending to and incorporating the genotype information in their treatment recommendations. Conversely, psychiatrists with more than 15 years of experience were not likely (with a membership probability close to 0) to consider the genotype information in their recommendations. To conclude the exploration of ANA in our data set, it is possible to use the same model with additional classes. More precisely, the full model requires creation of 16 separate classes to account for all possible patterns of ANA. Estimates from this model (not included in the article but

Table 1 – Model estimations for MNL, RPL, and RPL-ANA models.

Variable	MNL model		RPL model		RPL-ANA model	
	Estimate	SE	Estimate	SE	Estimate	SE
Change score	0.30*	0.02	0.44*	0.04	0.44*	0.04
σ change score			0.19*	0.03	0.20*	0.03
Genotype	-0.17†	0.09	-0.25	0.17	-2.02*	0.36
σ genotype			1.02*	0.17	0.16	0.91
Days	-0.08*	0.01	-0.11*	0.01	-0.11*	0.01
σ days			0.04*	0.02	0.04*	0.02
Weight gain	-0.08*	0.01	-0.11*	0.01	-0.10*	0.01
σ weight gain			0.04*	0.01	0.05*	0.01
% of psychiatrists who considered patient's genotype					15.6%	
% of psychiatrists who did not consider patient's genotype					84.4%	
Variation in ANA genotype info per year of experience					0.17*	0.06
Log likelihood	-594.69		-533.07		-532.63	
Parameters	4		8		9	

MNL, multinomial logit; RPL, random-parameters logit; RPL-ANA, random-parameters logit attribute nonattendance; SE, standard error.

* $P < 0.01$.

† $P < 0.05$.

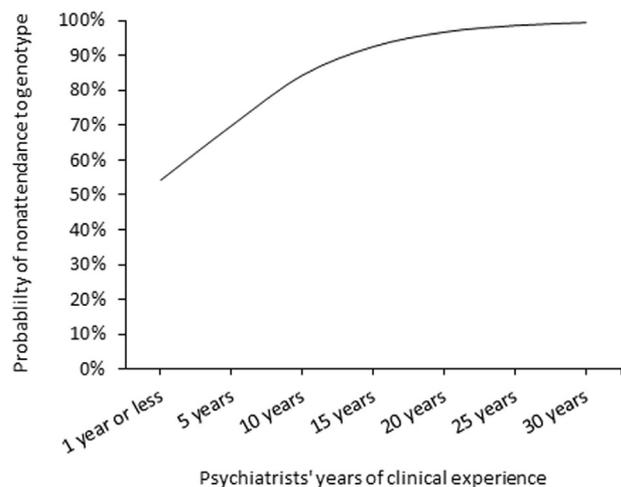


Fig. 2 – Psychiatrists' years of experience plotted against the probability of attending to the patient's genotype information. More experienced psychiatrists were less likely to attend to the genetic information of the patient.

available on request) suggest that only three classes have a membership probability significantly different from 0: full attendance (with a membership probability of 21.6%), nonattendance to genotype (60.5%), and nonattendance to both genotype and weight gain (with the lowest probability <10% and statistically insignificant). The membership probability of the last class, however, is reduced to almost 0 when the specification accounts for preference heterogeneity as in our model (RPL-ANA model) in Table 1. This makes the ANA specification proposed in Table 1 the most suitable to model the data from this study.

Discussion

This study investigated whether psychiatrists' treatment decisions are influenced by information about a patient's genotype even when they already know the patient's actual response to treatment. We provided psychiatrists with pre- and post-treatment patient outcomes, which identify a treatment's effectiveness, and information about the patients' genotype. Our premise was that the presence of a hyper-responsive genotype should not have influenced the treatments recommended by psychiatrists. Results suggested that most psychiatrists, but not all, were not influenced by the irrelevant genetic information about the patient. Years of clinical experience strongly determined whether psychiatrists incorporated the genetic information into their recommendations. Psychiatrists with 1 year or less of clinical experience had a 46% probability of responding to the genetic information. Psychiatrists with at least 15 years of experience had a 7% probability of incorporating the same genetic information.

Why were inexperienced psychiatrists more likely to be influenced by irrelevant genetic information about a patient? One possibility is that the prescribing behaviors of psychiatrists have undergone a gradual change over time, creating generational differences in their recommendations [37]. Another possibility is that, unlike more experienced practitioners, novice practitioners have been exposed to new discoveries in genetics and the potential value of patient genotype information as part of their medical training. Modern medical training has incorporated recent advances in genetics that were not known during the training of more experienced practitioners [38]. Nevertheless,

present medical training may not provide adequate guidance on when genetic information about a patient should be used and how it should be incorporated into clinical recommendations and prescriptions. For instance, in 2010, only 56% of a sample of 217 chief psychiatrists in psychiatric residency programs in the United States reported receiving training on genetics during their residency, and those who did received no more than 3 hours of training [39]. Thus, although novice psychiatrists may receive training on psychiatric genomics, directing their attention to its relevance in clinical practice, they may not receive sufficient training on the appropriate use of such patient information. We tentatively recommend that researchers and policymakers investigate more closely present education practices in terms of psychiatric genomics.

Our findings resonate with recent discoveries that clinicians' treatment recommendations can be influenced by subjective factors about a patient. For example, researchers have found that clinicians are less likely to recommend amniocentesis—an invasive prenatal test for genetic and chromosomal abnormalities—when pregnancies were conceived by assisted reproductive technologies than when they were conceived spontaneously, even though the method of conception is irrelevant to the possibility of genetic or chromosomal abnormalities [40]. Our present findings reveal that genetic information about a patient may also influence psychiatrists' treatment recommendations even when a patient's actual response to treatment is known, although this is less likely among experienced psychiatrists.

Our study has some limitations. We focused on the treatment recommendations of practicing psychiatrists. Further research is essential to assess how clinicians in other medical domains may be inappropriately influenced by genetic information in their medical decision making. In addition, we presented psychiatrists with hypothetical patient outcomes for hypothetical treatments rather than use actual patient outcomes for real treatments. We did so to control for potential redundancies between attributes and to allow a broad range of attribute levels. Studies have validated the use of vignettes to study individual preferences [41,42]. Nevertheless, the decisions in vignette-based studies usually do not have the same financial, psychosocial, or emotional consequences of treatment decisions made in clinical practice.

Conclusions

Building on encouraging results from past research on ANA in environmental economics [21,43,44] and health [14,45], our study confirms that ANA is a valuable tool for analyzing clinical decision making. To the authors' knowledge, this study is the first to suggest that less experienced psychiatrists may be inappropriately influenced by a patient's genetic information in their clinical decision making. Several authors have warned clinicians about being unduly influenced by a patient's genetic information, and it is plausible that more experienced clinicians may be more immune to the influence of a patient's genetic profile [5,46]. The findings of this study show that less experienced psychiatrists may be more susceptible to a form of pharmacogenetic exceptionalism, giving undue weight to a patient's genotype when they already know the patient's actual response to treatment. As a result, it is possible that less experienced psychiatrists will be less likely to recommend effective treatments or continue with ineffective treatment plans when they are aware of a patient's genetic profile.

We believe that the results of our present study may have important implications for medical practice. With the increased knowledge and awareness of the role that genes play in a patient's potential response to treatment, it is essential that

psychiatrists and other clinicians weigh this information appropriately in their clinical decision making. Understanding the role that genetics plays in treatment response could help clinicians maximize treatment response and minimize treatment side effects [2]. There is, however, a risk that too much weight could be given to a patient's genotype (pharmacogenetic exceptionalism) [5]. Psychiatrists and other health care professionals should be aware of the potential influence of a patient's genetic information on their clinical decision making, and this should be considered and highlighted during their education and further training.

Source of financial support: Financial support for this study was provided by a grant from the Department of Education and Learning. The funding agreement ensured the authors' independence in designing the study, interpreting the data, and writing and publishing the report.

REFERENCES

- [1] Jacob I, Awada AH, Payne K, Annemans L. Stratified medicine: a call for action. *Expert Rev Pharmacoecon Outcomes Res* 2013;13:277–9.
- [2] Trusheim MR, Berndt ER, Douglas FL. Stratified medicine: strategic and economic implications of combining drugs and clinical biomarkers. *Nat Rev Drug Discov* 2007;6:287–93.
- [3] Ikeda M, Tomita Y, Mouri A, et al. Identification of novel candidate genes for treatment response to risperidone and susceptibility for schizophrenia: integrated analysis among pharmacogenomics, mouse expression, and genetic case-control association approaches. *Biol Psychiatry* 2010;67:263–9.
- [4] Mikkelsen JD, Thomsen MS, Hansen HH, Lichota J. Use of biomarkers in the discovery of novel anti-schizophrenia drugs. *Drug Discov Today* 2010;15:137–41.
- [5] Kitsios GD, Kent DM. Personalised medicine: not just in our genes. *BMJ* 2012;344:e2161.
- [6] Johnson FR, Van Houtven G, Ozdemir S, et al. Multiple sclerosis patients' benefit-risk preferences: serious adverse event risks versus treatment efficacy. *J Neurol* 2009;256:554–62.
- [7] Johnson FR, Hauber B, Ozdemir S, et al. Are gastroenterologists less tolerant of treatment risks than patients? Benefit-risk preferences in Crohn's disease management. *J Manag Care Pharm* 2010;16:616–28.
- [8] Hauber A, Mohamed A, Johnson F. Quantifying asthma patient preferences for onset of effect of combination inhaled corticosteroids and long-acting beta 2-agonist maintenance medications. *Allergy Asthma Proc* 2009;30:139–47.
- [9] Bridges JFP, Mohamed AF, Finnern HW, et al. Patients' preferences for treatment outcomes for advanced non-small cell lung cancer: a conjoint analysis. *Lung Cancer* 2012;77:224–31.
- [10] De Bekker-Grob EW, Ryan M, Gerard K. Discrete choice experiments in health economics: a review of the literature. *Health Econ* 2012;21:145–72.
- [11] Food and Drug Administration. Patient preference information—submission, review in PMAs, HDE applications, and de novo requests, and inclusion in device labeling draft guidance for industry, Food and Drug Administration staff. 2015. Available from: <http://www.fda.gov/ucm/groups/fdagov-public/@fdagov-meddev-gen/documents/document/ucm446680.pdf> [Accessed May 6, 2016].
- [12] Gigerenzer G, Gaissmaier W. Heuristic decision making. *Annu Rev Psychol* 2011;62:451–82.
- [13] Campbell D, Hutchinson WG, Scarpa R. Incorporating discontinuous preferences into the analysis of discrete choice experiments. *Environ Resour Econ* 2008;41:401–17.
- [14] Erdem S, Campbell D, Thompson C. Elimination and selection by aspects in health choice experiments: prioritising health service innovations. *J Health Econ* 2014;38:10–22.
- [15] Hensher DA, Collins AT, Greene WH. Accounting for attribute non-attendance and common-metric aggregation in a probabilistic decision process mixed multinomial logit model: a warning on potential confounding. *Transportation (Amst)* 2012;40:1003–20.
- [16] Hensher DA, Rose JM. Simplifying choice through attribute preservation or non-attendance: implications for willingness to pay. *Transp Res E Logistics Transp Rev* 2009;45E:4:583–90.
- [17] Hensher D, Rose J, Greene W. The implications on willingness to pay of respondents ignoring specific attributes. *Transportation (Amst)* 2005;32:203–22.
- [18] Campbell D, Hensher D, Scarpa R. Cost thresholds, cut-offs and sensitivities in stated choice analysis: identification and implications. *Resour Energy Econ* 2012;34:396–411.
- [19] Scarpa R, Zanolli R, Bruschi V, Naspetti S. Inferred and stated attribute non-attendance in food choice experiments. *Am J Agric Econ* 2013;95:165–80.
- [20] Thieme M, Scarpa R, Galletto L, Boatto V. Sparkling wine choice from supermarket shelves: the impact of certification of origin and production practices. *Agric Econ* 2013;44:523–36.
- [21] Scarpa R, Gilbride TJ, Campbell D, Hensher DA. Modelling attribute non-attendance in choice experiments for rural landscape valuation. *Eur Rev Agric Econ* 2009;36:151–74.
- [22] Collins A. Attribute nonattendance in discrete choice models: measurement of bias, and a model for the inference of both nonattendance and taste heterogeneity. 2012. Available from: <http://ses.library.usyd.edu.au/handle/2123/8966>. [Accessed November 10, 2015].
- [23] Hole AR, Kolstad JR, Gyrd-Hansen D. Inferred vs. stated attribute non-attendance in choice experiments: a study of doctors' prescription behaviour. *J Econ Behav Organ* 2013;96:21–31.
- [24] Lagarde M. Investigating attribute non-attendance and its consequences in choice experiments with latent class models. *Health Econ* 2013;22:554–67.
- [25] Thurstone L. A law of comparative judgement. *Psychol Rev* 1927;34:273–86.
- [26] Manski C. The structure of random utility models. *Theory Decis* 1977;8:229–54.
- [27] McFadden D. Conditional logit analysis of qualitative choice behavior. In: Zarembka P, ed., *Frontiers in Econometrics*. New York, NY: Wiley, 1973;105–42.
- [28] Train K. *Discrete Choice Methods with Simulation* (2nd ed.). Cambridge, UK: Cambridge University Press, 2009.
- [29] McFadden D, Train K. Mixed [MNL] models for discrete response. *J Appl Econ* 2000;15:447–70.
- [30] Hensher D, Greene WH. The mixed logit model: the state of practice. *Transportation (Amst)* 2003;30:133–76.
- [31] Revelt D, Train K. Mixed logit with repeated choices. *Rev Econ Stat* 1998;80:647–57.
- [32] Bierlaire M. BIOGEME: a free package for the estimation of discrete choice models. Presented at: The 3rd Swiss Transport Research Conference. Ascona, Switzerland: Monte Verita, 2003.
- [33] Kay SR, Qjper LA. The Positive and Negative Syndrome Scale (PANSS) for schizophrenia. *Schizophr Bull* 1987;13:261.
- [34] Fenton WS, Chavez MR. Medication-induced weight gain and dyslipidemia in patients with schizophrenia. *Am J Psychiatry* 2006;163:1697–704.
- [35] Allison DB, Mentore JL, Heo M, et al. Antipsychotic-induced weight gain: a comprehensive research synthesis. *Am J Psychiatry* 1999;156:1686–96.
- [36] Ascher-Svanum H, Stensland M, Zhao Z, Kinon BJ. Acute weight gain, gender, and therapeutic response to antipsychotics in the treatment of patients with schizophrenia. *BMC Psychiatry* 2005;5:3.
- [37] Lukasiewicz M, Gasquet I, Casadebaig F, et al. Predictive factors of the number and the dose of anti-psychotics in a cohort of schizophrenic patients. *Pharmacoepidemiol Drug Saf* 2006;15:594–601.
- [38] Rubin E, Zorumski C. Psychiatric education in an era of rapidly occurring scientific advances. *Acad Med* 2003;78:351–4.
- [39] Winner J, Goebert D, Matsu C, Mrazek D. Training in psychiatric genomics during residency: a new challenge. *Acad Psychiatry* 2010;34:115–8.
- [40] Srebnik N, Miron-Shatz T, Rolison JJ, et al. Physician recommendation for invasive prenatal testing: the case of the "precious baby." *Hum Reprod* 2013;28:3007–11.
- [41] Luck J, Glassman P, Dresselhaus TR, Lee M. Comparison of vignettes, standardized patients, and chart abstraction. *JAMA* 2014;283:1715–22.
- [42] Mohan D, Fischhoff B, Farris C, et al. Validating a vignette-based instrument to study physician decision making in trauma triage. *Med Decis Making* 2013;34:242–52.
- [43] Campbell D, Hensher DA, Scarpa R. Non-attendance to attributes in environmental choice analysis: a latent class specification. *J Environ Plan Manag* 2011;54:1061–76.
- [44] Carlsson F, Kataria M, Lampi E. Dealing with ignored attributes in choice experiments on valuation of Sweden's environmental quality objectives. *Environ Resour Econ* 2010;47:65–89.
- [45] Hole AR. A discrete choice model with endogenous attribute attendance. *Econ Lett* 2011;110:203–5.
- [46] McKinnon RA, Ward MB, Sorich MJ. A critical analysis of barriers to the clinical implementation of pharmacogenomics. *Ther Clin Risk Manag* 2007;3:75–9.